## Math 3280A 22-10-06

Review

· Cumulative distributions

 $F_{x(b)} = P\{X \leq b\}, b \in \mathbb{R}$ 

$$F_{X} \text{ is non-decreasing, right cts, } \lim_{b \to +\infty} F_{X}(b) = 1 \lim_{b \to -\infty} F_{X}(b) = 0$$

$$P\{X=b\} = F_{X}(b) \to F_{X}(b-).$$

· Continuous r.V.'s,

X is said to be (absolutely) continuous if 
$$\exists a$$
  
nonnegative function on IR such that  
 $P\{X \in B\} = \int_{B} f(x) dx$   
for any "measurable" set  $B \subset IR$ .  
The function  $f$  is called the probability density function  
of X, or simply pdf of X.

<u>Remark</u>: In the continuous case,  $P\{X=a\}=o$  for any  $a\in \mathbb{R}$ .

Hence  $P\{a \le X \le b\} = P\{a < X \le b\} = P\{a \le X < b\} = P\{a < X < b\}$ 

Exer. 1. Suppose X has a Pdf  

$$f(x) = \begin{cases} \lambda e^{-\frac{x}{100}} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$
(1) Find the value of  $\lambda$ .  
(2) Find P { X > 100 }.  
Solution:  

$$1 = \int_{-\infty}^{\infty} f(x) dx$$

$$= \int_{0}^{\infty} \lambda e^{-\frac{x}{100}} dx$$

$$= \lambda \cdot (-100 \cdot e^{-\frac{x}{100}}) \Big|_{0}^{\infty}$$

$$= \lambda \cdot 100$$
Hence  $\lambda = \frac{1}{100}$ .

$$P\{X > 100\} = \int_{100}^{+\infty} f(x) dx$$

$$= \int_{100}^{+\infty} \frac{-x}{100} dx$$

$$= -e^{-\frac{x}{100}} \int_{100}^{\infty} \frac{-x}{100} dx$$

$$= -e^{-\frac{x}{100}} \int_{100}^{\infty} \frac{-x}{100} dx$$

$$= -e^{-1} \frac{-1}{100} \int_{100}^{\infty} \frac{-1}{100} dx$$
§ 5.2 Expectation of a discrete r.u. X is defined by  

$$E[X] = \sum x p\{X=x\},$$
where the summation is taken over all the possible values that  
X can take on.  
We can not directly use this method to define the expectation of  
a cts r.u., since in the cts case, X will take on uncountably  
diff values, and moreover,  

$$P\{X=x\}=0 \text{ for all } x \in \mathbb{R}.$$

Def. Let X be a cts r.v. with pdf f.  
Then we define  

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx.$$
Intuitive idea:  
In the continuous core, we make a partition  
of (-∞, ∞) by (X<sub>n</sub>)<sub>n=-∞</sub> Such that  
X<sub>n+1</sub> - X<sub>n</sub> =  $\Delta x$ .  
Then  

$$\sum_{n} x_n \cdot p\{x_n \le x_{n+1}\}$$

$$= \sum_{n} x_n \int_{x_n}^{x_n + \Delta x} f(x) dx$$

$$\approx \sum_{n} x_n f(x_n) \cdot \Delta x \rightarrow \int_{-\infty}^{\infty} x f(x) dx$$
as  $\Delta x \rightarrow 0$ .

Example 2. X is said to be uniformly distributed  
on [0,1] if it has the following density  
$$f(x) = \begin{cases} 1 & \text{if } x \in [0,1] \\ 0 & \text{otherwise} \end{cases}$$
  
Find E[X].  
Solution:  
 $E[X] = \int_{-\infty}^{\infty} x f(x) dx$   
 $= \int_{0}^{1} x \cdot 1 dx$   
 $= \frac{x^{2}}{2} \int_{0}^{1} = \frac{1}{2}$ .  
Below we consider the expectation of functions  
of cts r.u's.

Prop 3. Let X be a cts r.v. with density f.  
Let 9: 
$$[R \rightarrow IR$$
. Then  
 $E[9(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$ .  
We prove the above prop only in the case that  $9 \ge 0$ .  
To this end, we first prove the following.  
Lem 4. Let Y be a non-negative cts r.v. Then  
 $E[Y] = \int_{0}^{\infty} P\{Y > y\} dy$ .  
To prove Lem 4, we need to use the following result:  
 $\int_{a}^{b} (\int_{c}^{d} f(x, y) dx) dy = \int_{c}^{d} (\int_{a}^{b} f(x, y) dy) dx$   
when f is non-negative. This is a special version  
of the Fubini Theorem.

$$\begin{array}{rcl} & \label{eq:product} & \mbox{Pf of Lem 4.} & \mbox{Let f be the density of Y.} \\ & \mbox{Sinu Y is non-negative,} & \mbox{f(x)=o for x \le 0.} \\ & \mbox{So ELY]} = \int_{-\infty}^{\infty} x \mbox{f(x)} \mbox{dx} = \int_{0}^{\infty} x \mbox{f(x)} \mbox{dx}. & \mbox{Observe} \\ & \mbox{So EY} = \int_{0}^{\infty} \left( \int_{y}^{\infty} f(x) \mbox{dx} \right) \mbox{dy} \\ & = \int_{0}^{\infty} \left( \int_{y}^{\infty} f(x) \mbox{dx} + \int_{y}^{y} \mathbf{0} \cdot \mathbf{f(x)} \mbox{dx} \right) \mbox{dy} \\ & = \int_{0}^{\infty} \left( \int_{0}^{\infty} g(x, y) \mbox{f(x)} \mbox{dx} \right) \mbox{dy} \\ & = \int_{0}^{\infty} \left( \int_{0}^{\infty} g(x, y) \mbox{f(x)} \mbox{dx} \right) \mbox{dy} \\ & \mbox{dwhere } g: (\mathbf{0}, \mathbf{\infty}) \times (\mathbf{0}, \mathbf{\infty}) \rightarrow \mbox{IR is defined by} \\ & \mbox{q(x,y)} = \begin{cases} 1 & \mbox{if } x > y \\ \mathbf{0} & \mbox{if } x \le y \end{cases} \\ & \mbox{if } x \le y \\ \hline & \mbox{if } x \le y \end{cases} \\ & \mbox{Eucline} \\ & \mbox{f(x)} \mbox{f(x)} \mbox{dy} \end{pmatrix} \mbox{dx} \\ & \mbox{dx} \end{aligned}$$

$$= \int_{0}^{\infty} f(x) \left( \int_{0}^{x} \frac{1}{1} dy + \int_{x}^{\infty} 0 dy \right) dx$$

$$= \int_{0}^{\infty} x f(x) dx$$

$$= E[Y]. \qquad \square$$

$$Proof of Prop. 3: By Lem 4,$$

$$E[g(X)] = \int_{0}^{\infty} P\{g(X) > y\} dy.$$

$$W_{n}'t_{e} = \{x \in \mathbb{R} : g(x) > y\}. Then$$

$$P\{g(X) > y\} = P\{x \in \mathbb{B}\} = \int_{\mathbb{B}} f(x) dx = \int_{\{x : g(x) > y\}} f(x) dx.$$
Hence
$$E[g(X)] = \int_{0}^{\infty} \int_{\{x : g(x) > y\}} f(x) dx dy$$

$$= \int_{0}^{\infty} \int_{\{x : g(x) > y\}} f(x) dx dy$$

$$where \quad \Re(x, y) = \begin{cases} 1 & \text{if } g(x) > y \\ 0 & \text{otherwise.} \end{cases}$$

So by the Fubini Thm,  

$$E[g(X)] = \int_{-\infty}^{\infty} \left( \int_{0}^{\infty} h(x,y) f(x) dy \right) dx$$

$$= \int_{-\infty}^{\infty} f(x) \left( \int_{0}^{\infty} h(x,y) dy + \int_{q(x)}^{\infty} h(x,y) dy \right)$$

$$= \int_{-\infty}^{\infty} f(x) \left( \int_{0}^{g(x)} 1 dy + \int_{q(x)}^{\infty} 0 dy \right) dx$$

$$= \int_{-\infty}^{\infty} f(x) g(x) dx.$$
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